

Technical Appendices and Supplementary Material

A Further Discussion of Related Works

A.1 Comparison to Gastpar et al. (2024)

The estimability setting studied in our paper was introduced by Gastpar, Nachum, Shafer, and Weinberger (2024). In Theorem 3 of their paper, they show a limitation on estimability (a learnability–estimability trade-off) for algorithm-dependent bounds that is fairly abstract and involves a total variation condition that might be hard to check in many cases. In contrast, Theorems 3.1 and 3.2 involve very concrete combinatorial and geometric conditions (VC dimension, orthogonal functions). Theorems 4 and 5 in their paper are more concrete, but they hold only for exactly orthogonal functions with strict algebraic structure (parity functions). In contrast, our Theorem 3.2 applies generally to any nearly-orthogonal function class (including classes that are exactly-orthogonal as a special case).

Unlike Gastpar et al. (2024), our work also presents positive results (Theorem 4.3 and Fact C.2), showing cases where generalization bounds for specific algorithms can be tight (even if, e.g., uniform convergence does not hold). The conceptual connections between estimability, stability and conditional variance appearing in those results was not present in Gastpar et al. (2024).

Finally, our techniques also differ from those of Gastpar et al. (2024). We use the Johnson–Lindenstrauss lemma, our technical lemma (Lemma H.1), and the duality of linear programming — expanding the arsenal of tools readily available for the study of estimability.

In summary, our work builds upon the foundation laid by Gastpar et al. (2024), but we make several important contributions that go beyond their results.

A.2 Stability

In Definitions 4.1 and 4.2, we formalize simple stability conditions that guarantee the existence of tight generalization bounds, as we show in Theorem 4.3. There are many definitions of stability in the literature, and it is important to appreciate that Theorem 4.3 makes a nontrivial conceptual contribution by identifying the “correct” notion of stability for understanding estimability.

Definitions 4.1 and 4.2 are similar to the definition of hypothesis stability and loss stability in Kearns and Ron (1999), Elisseeff, Evgeniou, and Pontil (2005), and Rogers and Wagner (1978). Lei, Jin, and Ying (2022) use another similar definition for stability and utilize it to derive generalization bounds for GD and SGD.

In contrast, our definitions of stability are also reminiscent of the replace-one stability in Bousquet and Elisseeff (2002), but as we explain in Section 4, our definitions overcome an important limitation present in their definition. In particular, the memorization algorithm (Example 1.6), which is very estimable, is not stable according to the definition of stability of Bousquet and Elisseeff (2002), but it is stable according to our definitions.

A.3 Neural Tangent Kernel and Mean-Field Theory

There are many works that study generalization using the neural tangent kernel (NTK) or mean-field theory (MFT) approach.¹² To the best of our knowledge, these works do not provide general necessary or sufficient conditions for generalization bounds to be tight, which is the focus of our work. Additionally, they study generalization bounds for fairly specific families of algorithms such as gradient descent (or idealized versions thereof), while our work applies to a broader and more general class of algorithms.¹³

¹²E.g., Aminian et al. (2023), Chen et al. (2020), Nishikawa et al. (2022), and Nitanda et al. (2021).

¹³We note that Theorems 3.1 and 3.2 apply to learning algorithms that achieve 0 training error. Because this property is satisfied by many contemporary learning algorithms (even if the labels are random), we do not view this as a significant limitation on the generality of our results. This assumption is not essential, and it could easily be relaxed in future work.

B On Extending Our Results to Randomized Algorithms

For simplicity, in this paper we focus on deterministic learning rules. However, we recognize that the topic of randomized learning algorithms is very important, seeing as most algorithms used in practice today are randomized.

The estimability framework explored in this paper can be extended to handle randomized algorithms as well, and in fact the original work of Gastpar et al. (2024) already contains some initial treatment of randomized algorithms.

We expect that the results presented in this paper can be extended to randomized algorithms, and that the essence of the results remains mostly unchanged.

The first step in such an extension would be to clearly define what estimability means for randomized learning algorithms. A definition that one might initially consider is one where the estimator knows the randomness used by the algorithm, and must output a number that is with high probability close to the true population loss of the randomized algorithm. This definition is not very interesting, because a setting in which the estimator knows the randomness used by the randomized algorithm is equivalent to the setting of a deterministic algorithm, which is already covered by the results in this paper. Nonetheless, it is good to keep this definition in mind, because it means that our results for deterministic algorithms already apply as-is to randomized algorithms (like SGD) once the randomly chosen seed is fixed, which might be a simple and satisfactory approach for many purposes (SGD with a fixed random seed typically performs as well for most purposes as SGD with a fresh randomly-chosen seed).

Perhaps the more “correct” and interesting definition of estimability for randomized learning algorithms is one where the estimator knows the training set, but does not know the randomness used by the learning algorithm, and it is required to output a number that is close with high probability to the *expected* population loss of the randomized algorithm when executed with this training set (where the expectation is over the randomness of the algorithm). In this setting, we believe the essence of our results carries through, with an important conceptual difference: using randomness, one can always engineer a learning algorithm that is estimable, essentially by adding noise to the output of the algorithm. As the noise in the algorithm’s output increases, the expected 0-1 loss of the algorithm becomes closer to $1/2$, and so the algorithm becomes estimable with a trivial estimator that simply always outputs the number $1/2$. (With intermediate amounts of noise, a number between 0 and $1/2$ will be optimal).

Consequently, for randomized algorithms, our lower bounds in Theorems 3.1 and 3.2 can no longer be stated as absolute limitations on estimability. Rather there is now a trade-off between the performance of the algorithm and its estimability. As one adds more noise, the algorithm becomes more estimable, but its performance degrades. Thus, the corresponding theorems for randomized algorithm would state that no algorithm can simultaneously make good predictions for some large set of labeling functions and also be estimable.

On the other hand, the upper bound in Theorem 4.3 that states that stable algorithms are estimable remains basically unchanged for randomized algorithms.

To summarize, under a suitable definition of estimability for randomized algorithms, we expect that our results would not change much, though the statement (and proof) of the lower bounds would be somewhat more complex. We leave this work to future research.

C A Simple Characterization

The following definition is a variant of Definition 1.1. Such a variant allows us to have a simple characterization of estimability in Fact C.2. Namely, to understand whether an algorithm is estimable with respect to a set of distributions, one can examine the quantity $\mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [\text{var}(L_{\mathcal{D}}(A(S)) \mid S)]$.

Definition C.1. *Let \mathbb{D} be a set of distributions and let A be a learning algorithm. We say that A is (ε, m) -estimable in ℓ_2 with respect to \mathbb{D} , if there exists an estimator \mathcal{E} such that*

$$\mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [(\mathcal{E}(S) - L_{\mathcal{D}}(A(S)))^2] \leq \varepsilon$$

584 We remark that for bounded loss functions, one can move between Definition C.1 and
 585 Definition 1.1 using Markov's inequality. Furthermore, although the characterization in the
 586 following theorem is simple, it might provide a technical condition that will be useful for
 587 future work.

588 **Fact C.2.** *A is (ε, m) -estimable in ℓ_2 with respect to \mathbb{D} if and only if*

$$\mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [\text{var}(L_{\mathcal{D}}(A(S)) \mid S)] \leq \varepsilon.$$

589 *Proof of Fact C.2.* The result that the minimum mean-square error (MMSE) estimator
 590 corresponds to the conditional expectation is a well-established theorem in probability theory
 591 (see, for instance, Section 7.9 in Grimmett and Stirzaker (2020)). For the sake of completeness,
 592 we present a proof of this result.

593 We will use the following simple claim.

594 **Claim C.3.** *Let $c_1, \dots, c_k, p_1, \dots, p_k \in \mathbb{R}$ such that $\sum_{i=1}^k p_i = 1$, then*

$$\operatorname{argmin}_{x \in \mathbb{R}} \sum_{i=1}^k p_i \cdot (x - c_i)^2 = \sum_{i=1}^k p_i \cdot c_i.$$

595 The claim follows by taking the derivative of $\sum_{i=1}^k p_i \cdot (x - c_i)^2$ with respect to x which
 596 yields the equation:

597 $\sum_{i=1}^k 2p_i(x - c_i) = 0$ that implies $x = \sum_{i=1}^k p_i c_i$ since $\sum_{i=1}^k p_i = 1$.

598 The following shows that the estimator $\mathcal{E}^*(S) := \mathbb{E}[L_{\mathcal{D}}(A(S)) \mid S]$ is optimal and the
 599 inequality follows from Claim C.3. Let \mathcal{E} be any estimator for A .

$$\begin{aligned} & \mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [(\mathcal{E}(S) - L_{\mathcal{D}}(A(S)))^2] \\ &= \sum_S \mathbb{P}(S) \sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) (L_{\mathcal{D}}(A(S)) - \mathcal{E}(S))^2 \\ &= \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) (L_{\mathcal{D}}(A(S)) - \mathcal{E}(S))^2 \right] \\ &\geq \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) \left(L_{\mathcal{D}}(A(S)) - \sum_{\mathcal{D} \in \mathbb{D}} [\mathbb{P}(\mathcal{D} \mid S) L_{\mathcal{D}}(A(S))] \right)^2 \right] \\ &= \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) (L_{\mathcal{D}}(A(S)) - \mathcal{E}^*(S))^2 \right] \\ &= \mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [(\mathcal{E}^*(S) - L_{\mathcal{D}}(A(S)))^2]. \end{aligned}$$

600 This means that A is square loss (ε, m) -estimable with respect to \mathbb{D} if and only if \mathcal{E}^* can
 601 achieve ε accuracy. It achieves such accuracy if and only if $\mathbb{E}[\text{var}(L_{\mathcal{D}}(A(S)) \mid S)] \leq \varepsilon$. This
 602 follows by the following equalities that complete the proof.

$$\begin{aligned}
\mathbb{E}[\text{var}(L_{\mathcal{D}}(A(S)) \mid S)] &= \mathbb{E} \left[\mathbb{E} \left[(L_{\mathcal{D}}(A(S)) - \mathbb{E}[L_{\mathcal{D}}(A(S)) \mid S])^2 \mid S \right] \right] \\
&= \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) (L_{\mathcal{D}}(A(S)) - \mathbb{E}[L_{\mathcal{D}}(A(S)) \mid S])^2 \right] \\
&= \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) \left(L_{\mathcal{D}}(A(S)) - \sum_{\mathcal{D} \in \mathbb{D}} [\mathbb{P}(\mathcal{D} \mid S) L_{\mathcal{D}}(A(S))] \right)^2 \right] \\
&= \mathbb{E} \left[\sum_{\mathcal{D} \in \mathbb{D}} \mathbb{P}(\mathcal{D} \mid S) (L_{\mathcal{D}}(A(S)) - \mathcal{E}^*(S))^2 \right] \\
&= \mathbb{E}_{\mathcal{D} \sim \mathcal{U}(\mathbb{D}), S \sim \mathcal{D}^m} [(\mathcal{E}^*(S) - L_{\mathcal{D}}(A(S)))^2] \quad \square
\end{aligned}$$

603 D Details for Example 1.8

604 For sample size $m \geq d + 10$, any ERM algorithm for \mathcal{H} satisfies $p(m) \geq 0.999$, meaning it
605 learns \mathbb{D} well, and hence is $(0, 10^{-3}, d + 10)$ -estimable on average. This holds because for an
606 ERM to output the ground truth, it is clearly sufficient that only a single sample-consistent
607 function exists in the concept class (the ground truth). Similarly, in the event that there
608 are $t > 1$ sample-consistent functions, the success probability is given by $1/t$ due to the
609 uniform prior over ground truth distributions. Parity functions are fully characterized by
610 their coefficient vector $w = [w_1, \dots, w_d]$. Since the labels y are a bilinear function in the
611 inputs x and coefficients w , one can obtain w from $m \geq d$ linearly independent samples x_i
612 by solving the linear system of equation $y = Xw$ with design matrix $X \in \{0, 1\}^{m \times d}$. More
613 generally, X having rank $d - k$ is equivalent to the event of having $t = 2^k$ sample-consistent
614 functions (coefficient vectors) since every additional linearly independent row rules out half
615 of all parity functions. Now assume X consists of all i.i.d. $\text{Ber}(\frac{1}{2})$ entries and y contains
616 the labels of all samples. The probability of zero population loss can now be obtained from
617 the law of total probability with the probabilities of rank deficiency computed according to
618 Corollary 2.2 in Blake and Studholme (2006).

619 Similar calculations show that for smaller sample sizes, any ERM for \mathcal{H} satisfies $p(d) \geq 0.61$,
620 and $p(d - 1) \geq 0.38$. An application of Theorem 5 in Gastpar et al. (2024) shows that there
621 exist ERM algorithms such that for any $6 \leq m \leq d$ there exists a collection \mathbb{D}_m for which the
622 algorithm is not $(0.25, 0.32, m)$ -estimable on average. These algorithms have an inductive
623 bias towards a subset $\mathcal{F} \subseteq \mathcal{H}$, such that they perform well for distributions labeled by a
624 function from \mathcal{F} , and perform poorly for target functions from the complement of \mathcal{F} .

625 E Proof of Theorem 3.1

626 Recall the definition of nearly-orthogonal functions (Definition 2.8). The proof of Theorem 3.1
627 uses a corollary of the Johnson–Lindenstrauss lemma (Theorem K.1), which states that
628 random vectors in a high dimensional space are nearly orthogonal, as follows.¹⁴

629 **Claim E.1.** *Let $\varepsilon \in (0, 1/2)$, and let $d, n \in \mathbb{N}$ such that*

$$n \leq \exp(d\varepsilon^2/54).$$

630 *Let $\mathcal{U} = \mathcal{U}(\{\pm 1\}^{[d]})$ be the uniform distribution over functions $[d] \rightarrow \{\pm 1\}$, and consider a*
631 *random sequence F of functions F_1, \dots, F_n sampled independently from \mathcal{U} . Then*

$$\mathbb{P}_{F \sim \mathcal{U}^n} [F \in \perp_{\varepsilon, [d]}] \geq 0.99.$$

632 *Proof of Claim E.1.* If $n = 1$ there is nothing to prove, so we assume $n \geq 2$. Let $R \sim$
633 $\mathcal{U}(\{\pm 1\}^{d \times n})$ be a $d \times n$ matrix with entries in $\{\pm 1\}$ chosen independently and uniformly
634 at random. In particular, for each $i \in [n]$, the i -th column of R is a vector of d numbers in
635 $\{\pm 1\}$ chosen independently and uniformly at random. Hence, using e_1, \dots, e_n to denote the

¹⁴It is also possible to prove a similar claim by directly using concentration of measure (e.g., Hoeffding's inequality), without using the Johnson–Lindenstrauss lemma.

standard basis of \mathbb{R}^n , we identify the vector Re_i , which is the i -th column of R , with the random function $F_i : [d] \rightarrow \{\pm 1\}$.

Recall that for vectors $u, v \in \mathbb{R}^d$,

$$\|u - v\|_2^2 = \langle u - v, u - v \rangle = \|u\|_2^2 - 2\langle u, v \rangle + \|v\|_2^2,$$

so

$$\langle u, v \rangle = \frac{\|u\|_2^2 + \|v\|_2^2 - \|u - v\|_2^2}{2}. \quad (5)$$

Invoking Theorem K.1 with $s = n$, $\beta = 7$, $V = \{e_1, \dots, e_n\} \subseteq \mathbb{R}^n$, and d, n, ε as in the claim statement implies that

$$\mathbb{P}_{R \sim \mathcal{U}(\{\pm 1\}^{d \times n})} \left[\begin{array}{l} \forall i, j \in [n], i \neq j : \\ (1 - \varepsilon) \cdot 2 \leq \left\| \frac{1}{\sqrt{d}} Re_i - \frac{1}{\sqrt{d}} Re_j \right\|_2^2 \leq (1 + \varepsilon) \cdot 2 \end{array} \right] \geq 1 - \frac{1}{n^\beta}. \quad (6)$$

Hence, with probability at least $1 - 1/n^\beta \geq 1 - 1/2^7 \geq 0.99$ over the choice of F , every distinct $i, j \in [n]$ satisfy

$$\begin{aligned} |\mathbb{E}_{x \sim \mathcal{U}([d])} [F_i(x) F_j(x)]| &= \left| \frac{1}{d} \sum_{x \in [d]} F_i(x) F_j(x) \right| = \left| \frac{1}{d} \langle Re_i, Re_j \rangle \right| \quad (\text{Identifying } F_i \text{ with } Re_i) \\ &= \left| \frac{\|Re_i\|_2^2 + \|Re_j\|_2^2 - \|Re_i - Re_j\|_2^2}{2d} \right| \quad (\text{By Eq. (5)}) \\ &= \left| 1 - \frac{1}{2} \left\| \frac{1}{\sqrt{d}} Re_i - \frac{1}{\sqrt{d}} Re_j \right\|_2^2 \right| \\ &\leq \varepsilon, \quad (\text{By Eq. (6)}) \end{aligned}$$

as desired. \square

Proof of Theorem 3.1. Fix an \mathcal{H} -shattered set $\mathcal{X}_d \subseteq \mathcal{X}$ with cardinality $|\mathcal{X}_d| = d$, and for each $f : \mathcal{X}_d \rightarrow \{\pm 1\}$ let $\mathcal{D}_f = \mathcal{U}(\{(x, f(x)) : x \in \mathcal{X}_d\})$. Note that the distributions \mathcal{D}_f are \mathcal{H} -realizable. We will show that there exists a collection $\mathbb{D} = \{\mathcal{D}_f : f \in \mathcal{F}\}$ that satisfies Eq. (4), where $\mathcal{F} \subseteq \{\pm 1\}^{\mathcal{X}_d}$ is a set of $k = 2^m + 1$ functions.

Consider the following experiment:

1. Sample a sequence of functions $G = (G_1, \dots, G_k)$ independently and uniformly at random from $\{\pm 1\}^{\mathcal{X}_d}$.
2. Sample a function F uniformly from G .
3. Sample a sequence of points $X = (X_1, \dots, X_m)$ independently and uniformly at random from \mathcal{X}_d . (X is sampled independently of (G, F) .)
4. For each $i \in [m]$, let $Y_i = F(X_i)$, let $Y = (Y_1, \dots, Y_m)$, and let $S = ((X_1, Y_1), \dots, (X_m, Y_m))$.

Let \mathcal{P} be the joint distribution of (G, F, X, Y, S) . Consider the following events:

- $\mathcal{E}_1 = \{G \in \perp_{\varepsilon, \mathcal{X}_d}\}$ for $\varepsilon = 2/d^{1/4}$. By Claim E.1 and the choice of k , $\mathcal{P}(\mathcal{E}_1) \geq 0.99$ for d large enough.¹⁵
- $\mathcal{E}_2 = \{|\{X_1, \dots, X_m\}| = m\}$. By Claim K.2 and the choice of m , $\mathcal{P}(\mathcal{E}_2) \geq 0.99$.

¹⁵We choose $d_0 \in \mathbb{N}$ to be the universal constant such that this inequality holds for all integers $d \geq d_0$ and all $m \leq \sqrt{d}/10$.

661 • $\mathcal{E}_3 = \{|G_S| = 2\}$. $\mathcal{P}(\mathcal{E}_3 | \mathcal{E}_2) \geq 1/e$. To see this, note that each function $G_i \in G \setminus \{F\}$ is
 662 chosen independently of F . Hence, the probability that a function G_i agrees with F on
 663 the m distinct samples in X (i.e., the probability that $G_i(X_j) = F(X_j)$ for all $j \in [m]$,
 664 given \mathcal{E}_2) is $p = 2^{-m}$. The functions in G are chosen independently, so the number
 665 T of functions in $G \setminus \{F\}$ that agree with F on m distinct samples has a binomial
 666 distribution $T \sim \text{Bin}(k-1, p)$. So

$$\begin{aligned} \mathbb{P}[T = 1] &= (k-1) \cdot p \cdot (1-p)^{k-2} = (1-p)^{k-2} \\ &\geq \left(e^{-\frac{p}{1-p}}\right)^{k-2} & (\forall p < 1 : 1-p \geq e^{-p/(1-p)}) \\ &= 1/e. \end{aligned}$$

667 Let $\mathcal{E} = \mathcal{E}_1 \cap \mathcal{E}_3$. Combining the above bounds yields

$$\begin{aligned} \mathcal{P}(\mathcal{E}) &= \mathcal{P}(\mathcal{E}_1 \cap \mathcal{E}_3) \\ &\geq \mathcal{P}(\mathcal{E}_3) - \mathcal{P}(\mathcal{E}_1^C) \\ &\geq \mathcal{P}(\mathcal{E}_3 | \mathcal{E}_2) \cdot \mathcal{P}(\mathcal{E}_2) - \mathcal{P}(\mathcal{E}_1^C) \\ &\geq 0.99 \cdot 1/e - 0.01 > 1/3. \end{aligned}$$

668 By an averaging argument, this implies that there exists $\mathcal{F} \subseteq \{\pm 1\}^{\mathcal{X}_d}$ such that $\mathcal{F} \in \perp_{\varepsilon, \mathcal{X}_d}$
 669 for $\varepsilon = 2/d^{1/4}$ and

$$\mathcal{P}(|G_S| = 2 \mid G = \mathcal{F}) \geq 1/3. \quad (7)$$

670 Fix this \mathcal{F} , and let A be an \mathcal{F} -interpolating learning rule. From the technical lemma
 671 (Lemma H.1), there exists a collection of \mathcal{F} -realizable distributions $\mathbb{D} \subseteq \Delta(\mathcal{X}_d \times \{\pm 1\})$ such
 672 that for any estimator $\mathcal{E} : (\mathcal{X} \times \{\pm 1\})^m \rightarrow [0, 1]$ that may depend on \mathbb{D} and A ,

$$\begin{aligned} \mathbb{P}_{\substack{\mathcal{D} \sim \mathbb{U}(\mathbb{D}) \\ S \sim \mathcal{D}^m}} \left[|\mathcal{E}(S) - L_{\mathcal{D}}(A(S))| \geq \frac{1}{4} - \frac{\varepsilon}{4} \right] &\geq \frac{1}{2} \cdot \mathbb{P}_{\substack{\mathcal{D} \sim \mathbb{U}(\mathbb{D}) \\ S \sim \mathcal{D}^m}} [|\mathcal{F}_S| = 2] \\ &\geq \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{6}, \quad (\text{By Eq. (7)}) \end{aligned}$$

673 as desired. □

674 F Proof of Theorem 3.2

675 *Proof of Theorem 3.2.* We take $\mathbb{D} = \{\mathcal{D}_f : f \in \mathcal{F}\}$ where $\mathcal{D}_f = \mathbb{U}(\{(x, f(x)) : x \in \mathcal{X}\})$. Fix
 676 a function $f^* \in \mathcal{F}$, let $S \sim (\mathcal{D}_{f^*})^m$, and consider the random variable $Z = |\mathcal{F}_S|$. We bound
 677 the expectation and variance of Z , and then show a lower bound on the probability that
 678 $Z \in \{2, 3\}$.

679 Let $S = ((X_1, Y_1), \dots, (X_m, Y_m))$ and $X = \{X_1, \dots, X_m\}$, and let E denote the event in
 680 which $|X| = m$ (i.e., S is collision-free). For each $f \in \mathcal{F}$, let $Z_f = \mathbb{1}(\forall i \in [m] : f(X_i) = Y_i)$,
 681 so that $Z = \sum_{f \in \mathcal{F}} Z_f$.

$$\begin{aligned} \mathbb{E}_{S \sim (\mathcal{D}_{f^*})^m} [Z \mid E] &= \mathbb{E} \left[\sum_{f \in \mathcal{F}} Z_f \mid E \right] \\ &= 1 + \sum_{\substack{f \in \mathcal{F} \\ f \neq f^*}} \mathbb{P}[\forall i \in [m] : f(X_i) = Y_i \mid E] & (Z_{f^*} = 1) \\ &\leq 1 + 2^m \cdot \left(\frac{1}{2} + \frac{1}{2} \cdot \frac{1}{1000m} \right)^m & (\text{By Fact 2.9}) \\ &\leq 1 + e^{1/1000} < 2.002. & (8) \end{aligned}$$

$$\mathbb{E}_{S \sim (\mathcal{D}_{f^*})^m} [Z \mid E] \geq 1 + 2^m \cdot \left(\frac{1}{2} - \frac{1}{2} \cdot \frac{1}{1000m} \right)^m \quad (\text{By Fact 2.9})$$

$$\geq 1 + e^{-1/500}. \quad (1-x \geq e^{-x/(1-x)}) \quad (9)$$

$$\begin{aligned} \mathbb{E}_{S \sim (\mathcal{D}_{f^*})^m} [Z^2 \mid E] &= \mathbb{E} \left[\left(\sum_{f \in \mathcal{F}} Z_f \right) \left(\sum_{g \in \mathcal{F}} Z_g \right) \mid E \right] \\ &= \mathbb{E} \left[\left(1 + \sum_{\substack{f \in \mathcal{F} \\ f \neq f^*}} Z_f \right) \left(1 + \sum_{\substack{g \in \mathcal{F} \\ g \neq f^*}} Z_g \right) \mid E \right] \quad (Z_{f^*} = 1) \\ &= \mathbb{E} \left[1 + 2 \sum_{\substack{f \in \mathcal{F} \\ f \neq f^*}} Z_f + \sum_{\substack{f \in \mathcal{F} \\ f \neq f^*}} \sum_{\substack{g \in \mathcal{F} \\ g \neq f^*}} Z_f Z_g \mid E \right] \\ &= \mathbb{E} \left[1 + 3 \sum_{\substack{f \in \mathcal{F} \\ f \neq f^*}} Z_f + \sum_{\substack{f, g \in \mathcal{F} \setminus \{f^*\} \\ f \neq g}} Z_f Z_g \mid E \right] \\ &= 1 + 3(\mathbb{E}[Z \mid E] - 1) + \sum_{\substack{f, g \in \mathcal{F} \setminus \{f^*\} \\ f \neq g}} \mathbb{E}[Z_f Z_g \mid E]. \end{aligned} \quad (10)$$

$$\begin{aligned} \sum_{\substack{f, g \in \mathcal{F} \setminus \{f^*\} \\ f \neq g}} \mathbb{E}[Z_f Z_g \mid E] &= \sum_{\substack{f, g \in \mathcal{F} \setminus \{f^*\} \\ f \neq g}} \mathbb{P}[\forall i \in [m] : f(X_i) = g(X_i) = f^*(X_i) \mid E] \\ &\leq 2^{2m} \cdot \left(\frac{1}{4} + \frac{3}{4} \cdot \frac{1}{1000m} \right)^m \quad (\text{By Claim J.1}) \\ &= \left(1 + \frac{3}{1000m} \right)^m \\ &\leq e^{3/1000}. \end{aligned} \quad (11)$$

Combining Eqs. (8) to (11) yields

$$\begin{aligned} \text{Var}[Z \mid E] &= \mathbb{E}[Z^2 \mid E] - (\mathbb{E}[Z \mid E])^2 \\ &\leq 1 + 3e^{1/1000} + e^{3/1000} - \left(1 + e^{-1/500} \right)^2 \\ &< 1.02. \end{aligned}$$

By Lemma I.1,

$$\mathbb{P}[Z \in \{2, 3\} \mid E] \geq 1 - \frac{\text{Var}[Z \mid E]}{2} \geq 0.49.$$

Claim K.2 and $|\mathcal{X}'| \geq 100m^2$ imply that $\mathbb{P}[E] \geq 0.99$. Hence,

$$\mathbb{P}[Z \in \{2, 3\}] \geq \mathbb{P}[E] \cdot \mathbb{P}[Z \in \{2, 3\} \mid E] \geq 0.99 \cdot 0.49 \geq 0.48. \quad (12)$$

Finally, invoking our technical lemma (Lemma H.1) yields

$$\mathbb{P}_{\substack{F \sim \mathcal{U}(\mathcal{F}) \\ S \sim (\mathcal{D}_F)^m}} \left[|\mathcal{E}(S) - L_{\mathcal{D}_F}(A(S))| \geq \frac{1}{4} - \frac{1}{4000m} \right] \geq \frac{\mathbb{P}[Z \in \{2, 3\}]}{3} \geq 0.16,$$

as desired. \square

G Proof of Theorem 4.3

Proof. If (A, \mathbb{D}) is (α, β, m, k) -hypothesis stable, then in particular (A, \mathbb{D}) is also (α, β, m, k) -loss stable. Hence, it suffices to prove the claim for the case of loss stability. We construct a uniform estimator \mathcal{E} as follows. Given a sample $S \in \mathcal{Z}^m$ for $\mathcal{Z} = (\mathcal{X} \times \{\pm 1\})$, let $S_1 \circ S_2 = S$ be the partition of S such that $S_1 \in \mathcal{Z}^{m-k}$ and $S_2 \in \mathcal{Z}^k$. Take $\mathcal{E}(S) = L_{S_2}(A(S_1))$.

By the triangle inequality,

$$|\mathcal{E}(S) - L_{\mathcal{D}}(A(S))| \leq |\mathcal{E}(S) - L_{\mathcal{D}}(A(S_1))| + |L_{\mathcal{D}}(A(S_1)) - L_{\mathcal{D}}(A(S))|,$$

so

$$\begin{aligned} \mathbb{P}_{S \sim \mathcal{D}^m}[|\mathcal{E}(S) - L_{\mathcal{D}}(A(S))| > \varepsilon] &\leq \mathbb{P}\left[\begin{array}{l} |L_{S_2}(A(S_1)) - L_{\mathcal{D}}(A(S_1))| > \alpha_0 \vee \\ |L_{\mathcal{D}}(A(S_1)) - L_{\mathcal{D}}(A(S))| > \alpha_1 \end{array} \right] \\ &\leq \mathbb{P}[|L_{S_2}(A(S_1)) - L_{\mathcal{D}}(A(S_1))| > \alpha_0] \\ &\quad + \mathbb{P}[|L_{\mathcal{D}}(A(S_1)) - L_{\mathcal{D}}(A(S))| > \alpha_1] \\ &\leq \beta_0 + \beta_1 = \delta, \end{aligned}$$

where the final step follows from Hoeffding's inequality, the choice of k , and the stability of A . \square

H Technical Lemma for Inestimability

Lemma H.1. Let $m \in \mathbb{N}$, let $\varepsilon > 0$, let \mathcal{X} be a finite set, let $\mathcal{F} \subseteq \{\pm 1\}^{\mathcal{X}}$ such that $\mathcal{F} \in \perp_{\varepsilon, \mathcal{X}}$, and let $A : (\mathcal{X} \times \{\pm 1\})^m \rightarrow \{\pm 1\}^{\mathcal{X}}$ be an \mathcal{F} -interpolating learning rule. For each $f \in \mathcal{F}$ let $\mathcal{D}_f = \mathcal{U}(\{(x, f(x)) : x \in \mathcal{X}\})$, and for each $k \in \mathbb{N}$ let

$$p_k = \mathbb{P}_{\substack{F \sim \mathcal{U}(\mathcal{F}) \\ S \sim (\mathcal{D}_F)^m}}[|\mathcal{F}_S| = k].$$

Then for any estimator $\mathcal{E} : (\mathcal{X} \times \{\pm 1\})^m \rightarrow [0, 1]$ that may depend on A ,

$$\mathbb{P}_{\substack{F \sim \mathcal{U}(\mathcal{F}) \\ S \sim (\mathcal{D}_F)^m}} \left[|\mathcal{E}(S) - L_{\mathcal{D}_F}(A(S))| \geq \frac{1}{4} - \frac{\varepsilon}{4} \right] \geq \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \frac{p_k}{k}.$$

Proof. Consider the following experiment:

1. Sample a sequence of points $X = (X_1, \dots, X_m)$ independently and uniformly at random from \mathcal{X} .
2. Sample a function F uniformly from \mathcal{F} , independently of X .
3. For each $i \in [m]$, let $Y_i = F(X_i)$, let $Y = (Y_1, \dots, Y_m)$, and let $S = ((X_1, Y_1), \dots, (X_m, Y_m))$.

Let \mathcal{P} be the joint distribution of (X, F, Y, S) . Fix $k \in \{2, \dots, |\mathcal{F}|\}$, and let

$$s = ((x_1, y_1), \dots, (x_m, y_m)) \in (\mathcal{X} \times \{\pm 1\})^m$$

with $x = (x_1, \dots, x_m)$ and $y = (y_1, \dots, y_m)$ such that $|\mathcal{F}_s| = k$. Denote $\mathcal{F}_s = \{f_1, \dots, f_k\}$.

Then for any $i, j \in [k]$, $i \neq j$,

$$\begin{aligned} \mathcal{P}(S = s \mid F = f_i) &= \mathcal{P}(X = x \mid F = f_i) \\ &= \mathcal{P}(X = x \mid F = f_j) \quad (X \perp F) \\ &= \mathcal{P}(S = s \mid F = f_j). \end{aligned} \tag{13}$$

So,

$$\begin{aligned} \mathcal{P}(F = f_i \mid S = s) &= \frac{\mathcal{P}(S = s \mid F = f_i) \cdot \mathcal{P}(F = f_i)}{\mathcal{P}(S = s)} \\ &= \frac{\mathcal{P}(S = s \mid F = f_j) \cdot \mathcal{P}(F = f_j)}{\mathcal{P}(S = s)} \quad (\text{By Eq. (13), } F \sim \mathcal{U}(\mathcal{F})) \\ &= \mathcal{P}(F = f_j \mid S = s), \end{aligned} \tag{14}$$

711 Seeing as $\mathcal{P}(F \in \mathcal{F}_s \mid S = s) = 1$, this implies that for all $i \in [k]$, $\mathcal{P}(F = f_i \mid S = s) = 1/k$.
 712 Because A is \mathcal{F} -interpolating, $A(s) \in \mathcal{F}_s$. Without loss of generality, denote $A(s) = f_1$. From
 713 $\mathcal{F} \in \perp_{\varepsilon, \mathcal{X}}$ and Fact 2.9, $L_{\mathcal{D}_{f_i}}(f_j) \geq \frac{1}{2} - \frac{\varepsilon}{2} := 2\alpha$ for all $i, j \in [k], i \neq j$. Hence,

$$\begin{aligned} \mathcal{P}(L_{\mathcal{D}_F}(A(S)) = 0 \mid S = s) &= \mathcal{P}(F = A(S) \mid S = s) && (F, A(s) \in \mathcal{F}_s) \\ &= \mathcal{P}(F = f_1 \mid S = s) && (A(s) = f_1) \\ &= 1/k, \end{aligned} \tag{15}$$

714 and

$$\begin{aligned} \mathcal{P}(L_{\mathcal{D}_F}(A(S)) \geq 2\alpha \mid S = s) &= \mathcal{P}(F \neq A(S) \mid S = s) \\ &= \mathcal{P}(F \in \{f_2, \dots, f_k\} \mid S = s) \\ &= (k-1)/k. \end{aligned} \tag{16}$$

715 Hence, for any $\eta \in \mathbb{R}$,

$$\mathcal{P}(|L_{\mathcal{D}_F}(A(S)) - \eta| \geq \alpha \mid S = s) \geq \frac{1}{k}. \tag{17}$$

716 We conclude that for any estimator $\mathcal{E} : (\mathcal{X} \times \{\pm 1\})^m \rightarrow \mathbb{R}$,

$$\begin{aligned} &\mathcal{P}(|L_{\mathcal{D}_F}(A(S)) - \mathcal{E}(S)| \geq \alpha) \\ &\geq \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \mathcal{P}\left(|L_{\mathcal{D}_F}(A(S)) - \mathcal{E}(S)| \geq \alpha \bigwedge |\mathcal{F}_S| = k\right) \\ &= \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \sum_{s: |\mathcal{F}_s| = k} \mathcal{P}(|L_{\mathcal{D}_F}(A(S)) - \mathcal{E}(S)| \geq \alpha \mid S = s) \cdot \mathcal{P}(S = s) \\ &\geq \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \sum_{s: |\mathcal{F}_s| = k} \inf_{\eta \in \mathbb{R}} \mathcal{P}(|L_{\mathcal{D}_F}(A(S)) - \eta| \geq \alpha \mid S = s) \cdot \mathcal{P}(S = s) \\ &\geq \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \sum_{s: |\mathcal{F}_s| = k} \frac{1}{k} \cdot \mathcal{P}(S = s) && (\text{By Eq. (17)}) \\ &= \sum_{k \in \{2, \dots, |\mathcal{F}|\}} \frac{1}{k} \cdot \mathcal{P}(|\mathcal{F}_S| = k) \end{aligned}$$

717 as desired. □

718 I Concentration Bound via Linear Programming

719 **Lemma I.1.** *Let $n \in \mathbb{N}$, $v_{\max} \in \mathbb{R}$. Let Z be a random variable taking values in $[n]$ such*
 720 *that $\mu = \mathbb{E}[Z] \in [2, \sqrt{2} + 1]$ and $\text{Var}[Z] \leq v_{\max}$. Then $\mathbb{P}[Z \in \{2, 3\}] \geq 1 - v_{\max}/2$.*

721 We prove this concentration of measure bound using the duality of linear programs (see
 722 Section 7.4.1 in Boyd and Vandenberghe, 2014 for an exposition of this approach).

723 *Proof.* Let $Z' = Z - \mu$. Z' is a random variable with $\mathbb{E}[Z'] = 0$ and $\text{Var}[Z'] =$
 724 $\text{Var}[Z]$. Furthermore, $\mathbb{P}[Z \in \{2, 3\}] = \mathbb{P}[Z' \in \{2 - \mu, 3 - \mu\}]$. We show a lower bound
 725 on $\mathbb{P}[Z' \in \{2 - \mu, 3 - \mu\}]$ across all distribution of Z' with the above moment constraints.

726 Indeed, let X be a random variable taking values in $\{1 - \mu, 2 - \mu, \dots, n - \mu\}$ with $\mathbb{E}[X] = 0$ and
 727 $\text{Var}[X] \leq v_{\max}$ such that $\mathbb{P}[X \in \{2 - \mu, 3 - \mu\}]$ is minimal. In particular, the distribution of
 728 X is a solution to the following minimization problem.

$$\begin{aligned} &\min_{\mathcal{D}_X} \mathbb{P}[X \in \{2 - \mu, 3 - \mu\}] \\ &\text{s.t.} \\ &\quad \mathbb{E}[X] = 0 \\ &\quad \text{Var}[X] \leq v_{\max} \end{aligned}$$

729 The minimization problem can be formulated as a linear program with variables $p_k =$
730 $\mathbb{P}[X = k - \mu]$ for each $k \in [n]$.

$$\begin{aligned} & \min_{\mathcal{D}_X} p_2 + p_3 \\ & \text{s.t.} \\ & \sum_{k \in [n]} p_k \geq 1 \\ & \sum_{k \in [n]} -p_k \geq -1 \\ & \sum_{k \in [n]} p_k \cdot (k - \mu) \geq 0 \\ & \sum_{k \in [n]} p_k \cdot (\mu - k) \geq 0 \\ & \sum_{k \in [n]} -p_k \cdot (k - \mu)^2 \geq -v_{\max} \\ & \forall k \in [n] : p_k \geq 0. \end{aligned}$$

731 This linear program can be represented as

$$\begin{aligned} & \min (0, 1, 1, 0, \dots, 0) \cdot p \\ & \text{s.t.} \\ & \begin{pmatrix} 1 & 1 & \dots & 1 \\ -1 & -1 & \dots & -1 \\ 1 - \mu & 2 - \mu & \dots & n - \mu \\ \mu - 1 & \mu - 2 & \dots & \mu - n \\ -(1 - \mu)^2 & -(2 - \mu)^2 & \dots & -(n - \mu)^2 \end{pmatrix} \begin{pmatrix} p_1 \\ \vdots \\ p_n \end{pmatrix} \geq \begin{pmatrix} 1 \\ -1 \\ 0 \\ 0 \\ -v_{\max} \end{pmatrix} \\ & p \geq 0. \end{aligned}$$

732 Recall the symmetric duality

$$\begin{aligned} & \min c^T x \\ & \text{s.t.} \\ & Ax \geq b \\ & x \geq 0 \end{aligned} \quad \longleftrightarrow \quad \begin{aligned} & \max b^T y \\ & \text{s.t.} \\ & A^T y \leq c \\ & y \geq 0. \end{aligned}$$

734 Hence, the dual linear program is

$$\begin{aligned} & \max (1, -1, 0, 0, -v_{\max}) \cdot y \\ & \text{s.t.} \\ & \begin{pmatrix} 1 & -1 & 1 - \mu & \mu - 1 & -(1 - \mu)^2 \\ 1 & -1 & 2 - \mu & \mu - 2 & -(2 - \mu)^2 \\ 1 & -1 & 3 - \mu & \mu - 3 & -(3 - \mu)^2 \\ & & & \vdots & \\ 1 & -1 & n - \mu & \mu - n & -(n - \mu)^2 \end{pmatrix} \begin{pmatrix} y_1 \\ \vdots \\ y_5 \end{pmatrix} \leq \begin{pmatrix} 0 \\ 1 \\ 1 \\ 0 \\ 0 \end{pmatrix} \\ & y \geq 0. \end{aligned}$$

735 A direct calculation shows that the vector

$$y^* = (1, 0, \alpha, 0, \tfrac{1}{2}), \quad \alpha = \frac{1}{\mu - 1} - \frac{\mu - 1}{2}$$

736 is a feasible solution for the dual program for any $\mu \in [2, \sqrt{2} + 1]$. The value of the dual
737 program at y^* is $u = 1 - v_{\max}/2$. The weak duality theorem for linear programs implies that
738 u is a lower bound on the value of the primal problem. Hence,

$$\min \mathbb{P}[X \in \{2 - \mu, 3 - \mu\}] \geq u.$$

739 This implies that $\mathbb{P}[Z \in \{2, 3\}] \geq u$, as desired. \square

740 J Agreement Between Nearly-Orthogonal Functions

741 **Claim J.1.** *Let $\varepsilon > 0$, let \mathcal{X} be a set, and let $f, g, h : \mathcal{X} \rightarrow \{\pm 1\}$ such that $\{f, g, h\} \in \perp_{\varepsilon, \mathcal{X}}$.
742 Then $\mathbb{P}_{x \sim U(\mathcal{X})}[f(x) = g(x) = h(x)] \leq \frac{1}{4} + \frac{3\varepsilon}{4}$.*

743 *Proof.* Denote

$$\begin{aligned} a &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) = g(x) = h(x)] \\ b &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) \neq g(x) = h(x)] \\ c &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) = g(x) \neq h(x)] \\ d &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) \neq g(x) \neq h(x)] \end{aligned}$$

744 From $\{f, g, h\} \in \perp_{\varepsilon, \mathcal{X}}$ and Fact 2.9,

$$\begin{aligned} a + b &= \mathbb{P}_{x \sim U(\mathcal{X})}[g(x) = h(x)] \leq \frac{1}{2} + \frac{\varepsilon}{2} \\ a + c &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) = g(x)] \leq \frac{1}{2} + \frac{\varepsilon}{2} \\ a + d &= \mathbb{P}_{x \sim U(\mathcal{X})}[f(x) = h(x)] \leq \frac{1}{2} + \frac{\varepsilon}{2}. \end{aligned}$$

745 Adding these inequalities yields

$$3a + b + c + d \leq \frac{3}{2} + \frac{3\varepsilon}{2}.$$

746 From the identity $a + b + c + d = 1$,

$$2a \leq \frac{1}{2} + \frac{3\varepsilon}{2},$$

747 so $a \leq \frac{1}{4} + \frac{3\varepsilon}{4}$, as desired. □

748 K Miscellaneous Lemmas

749 The following result from Achlioptas (2003) is a variant of a lemma of Johnson and Linden-
750 Strauss (1984).

751 **Theorem K.1** (Johnson–Lindenstrauss). *Let $n, s \in \mathbb{N}$, let $\varepsilon, \beta > 0$, and let $V \subseteq \mathbb{R}^s$ be a set
752 with cardinality $|V| = n$. Let $d \in \mathbb{N}$ such that*

$$d \geq \frac{4 + 2\beta}{\varepsilon^2/2 - \varepsilon^3/3} \ln(n).$$

753 *Let R be a $d \times s$ random matrix such that each entry is chosen independently and uniformly
754 at random from $\{\pm 1\}$. Let $f_R : \mathbb{R}^s \rightarrow \mathbb{R}^d$ be given by $f_R(v) = (1/\sqrt{d}) \cdot Rv$. Then*

$$\mathbb{P}_{R \sim U(\{\pm 1\}^{d \times s})}[\forall u, v \in V : (1 - \varepsilon)\|u - v\|_2^2 \leq \|f_R(u) - f_R(v)\|_2^2 \leq (1 + \varepsilon)\|u - v\|_2^2] \geq 1 - \frac{1}{n^\beta}.$$

755 **Claim K.2** (Converse to Birthday Paradox). *Let $d, m \in \mathbb{N}$, and let $\beta \in (0, 1)$. If*

$$m \leq \min \left\{ \sqrt{d \ln \left(\frac{1}{\beta} \right)}, \frac{d}{2} \right\}$$

756 *then $\mathbb{P}_{X \sim (U([d]))^m}[|X| = m] \geq \beta$.*

757 *Proof.* We use the inequality $1 - x \geq e^{-x/(1-x)}$, which holds for $x < 1$.

$$\begin{aligned} \mathbb{P}_{X \sim (U([d]))^m}[|X| = m] &= 1 \cdot \left(1 - \frac{1}{d}\right) \cdot \left(1 - \frac{2}{d}\right) \cdots \left(1 - \frac{m-1}{d}\right) \\ &\geq \prod_{k=0}^{m-1} \exp\left(-\frac{k}{d-k}\right) = \exp\left(-\sum_{k=0}^{m-1} \frac{k}{d-k}\right) \\ &\stackrel{(*)}{\geq} \exp\left(-\frac{2}{d} \sum_{k=0}^{m-1} k\right) \geq \exp\left(-\frac{m^2}{d}\right), \end{aligned}$$

758 where $(*)$ follows from $m \leq d/2$. Solving $\exp\left(-\frac{m^2}{d}\right) \geq \beta$ yields the desired bound. \square

759 **Theorem K.3** (Hoeffding, 1963). *Let $a, b, \mu \in \mathbb{R}$ and $m \in \mathbb{N}$. Let Z_1, \dots, Z_m be a sequence*
760 *of i.i.d. real-valued random variables and let $Z = \frac{1}{m} \sum_{i=1}^m Z_i$. Assume that $\mathbb{E}[Z] = \mu$, and*
761 *for every $i \in [m]$, $\mathbb{P}[a \leq Z_i \leq b] = 1$. Then, for any $\varepsilon > 0$,*

$$\mathbb{P}[|Z - \mu| > \varepsilon] \leq 2 \exp\left(\frac{-2m\varepsilon^2}{(b-a)^2}\right).$$

762 L Experiments

763 L.1 Motivation and Setup

764 Here, we examine if there are practical algorithms that admit loss stability or even hypothesis
765 stability with substantial numerical values. To this end, we conduct experiments over a
766 simple neural network architecture across four datasets: MNIST, FashionMNIST, CIFAR10,
767 and CIFAR10 with random labels (figures 1-4, respectively). Throughout all experiments, we
768 employ one-hidden-layer perceptrons with 512 hidden neurons. We train the models using
769 stochastic gradient descent (SGD) with a momentum factor of 0.9 and a batch size of 1000,
770 optimizing the cross-entropy loss. For every data set, we train the models across learning
771 rates 0.1, 0.035,¹⁶ and 0.01. We average all the curves over 10 random seeds (tied for the
772 pairs of networks) and plot the standard deviation for all the curves.

773 The training procedure is as follows: we train two models in tandem, starting from the same
774 random initialization. The first model is provided with the full training set, whereas the
775 second model has $k = 100$ data points removed from its training set. These points are drawn
776 uniformly at random before the beginning of the training, and fixed thereafter. After each
777 epoch, we evaluate the training accuracy, test accuracy and hypothesis stability, i.e., the
778 agreement between the two models (which we calculate across the test set).

779 We set our main focus on the agreement of the models since the most amenable way to
780 show loss stability might be by way of proving hypothesis stability. The latter can perhaps
781 be mathematically proven in the case of neural networks by analyzing the stability of the
782 training dynamics under two slightly different training sets.

783 L.2 Results

784 Across all experiments, the training and test accuracy of the model pairs are essentially
785 identical throughout the training process. This suggests that at least simple models are loss
786 stable across vision tasks. In order to reduce visual clutter, we hence only plot training and
787 test accuracy of the first model (which has access to the full training set), respectively.

788 We observe higher agreement for simpler data sets and smaller learning rates. For example,
789 the learning rate has a considerable effect on agreement for CIFAR10 (≈ 0.65 for learning
790 rate 0.1 vs ≈ 0.8 for learning rate 0.01).

791 The key takeaway from Figures 1 through 4 is that the agreement is consistently higher than
792 the test accuracy. This relationship ensures that when applying the estimation procedure
793 outlined in Theorem 4.3, we can avoid vacuous predictions of perfect accuracy. In the
794 scenarios presented, the estimated accuracy will always be bounded away from 1, as it can
795 be expressed as *test error* + *(1 - agreement)*. For instance, with a learning rate of 0.01, the
796 maximum estimated accuracies are: 98% for MNIST (compared to 97.5% test accuracy),
797 90% for FashionMNIST (87% test accuracy), 72% for CIFAR10 (52% test accuracy), and
798 65% for CIFAR10 with random labels (10% test accuracy). These results illustrate a strong
799 correlation between stability estimation and data complexity.

800 We repeat the same experiments, modifying the width of the hidden layer to investigate its
801 impact on stability. The results, summarized in Table 1, reveal a strong positive correlation
802 between network width and stability. This effect is particularly pronounced for more complex
803 tasks, such as CIFAR10 and CIFAR10 with random labels. For instance, in the CIFAR10

¹⁶Except for CIFAR10, we present the results only for learning rate 0.1 and 0.01 to prevent clutter. The qualitative results are consistent across all datasets; that is, the curves of learning rate 0.035 lie between the curves of learning rate 0.1 and 0.01.

804 random labels setting with a learning rate of 0.01, increasing the width from 256 to 1024
805 neurons improves agreement from 32% to 50%, highlighting the stabilizing effect of greater
806 network width.

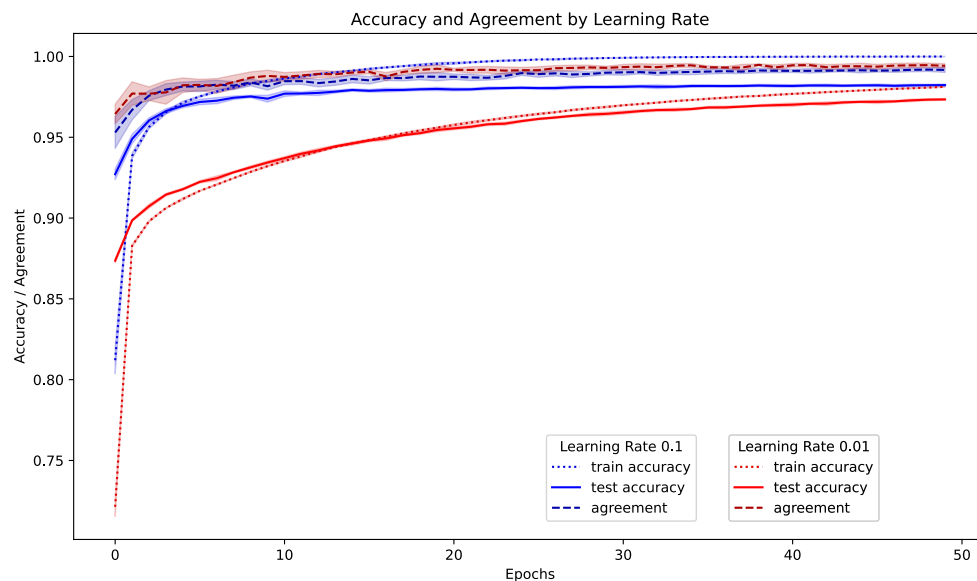


Figure 1: MNIST

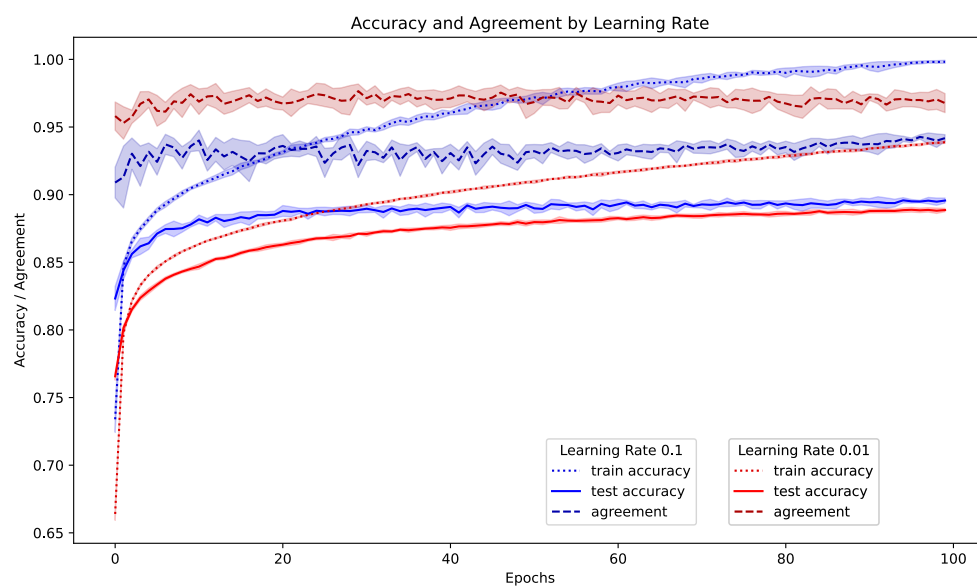


Figure 2: FashionMNIST

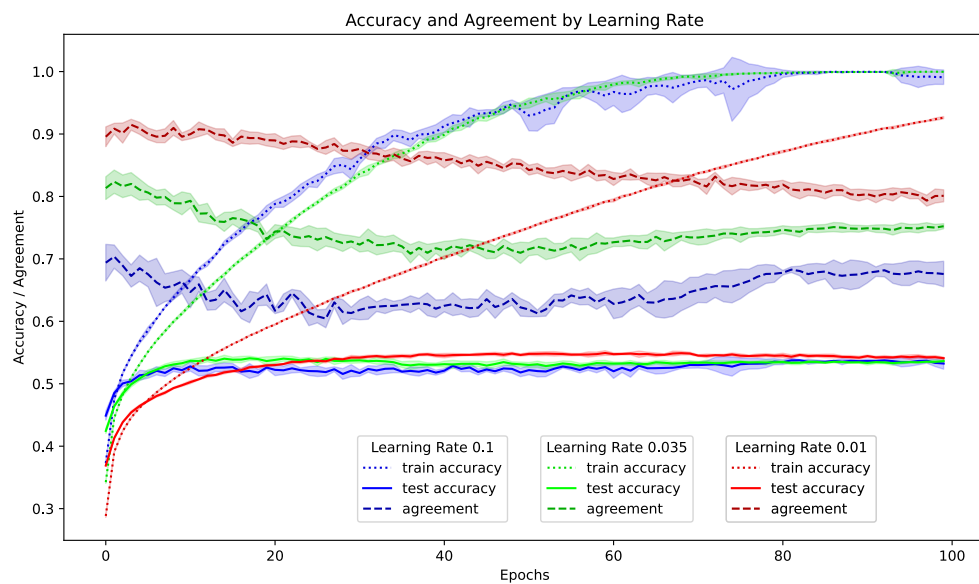


Figure 3: CIFAR10

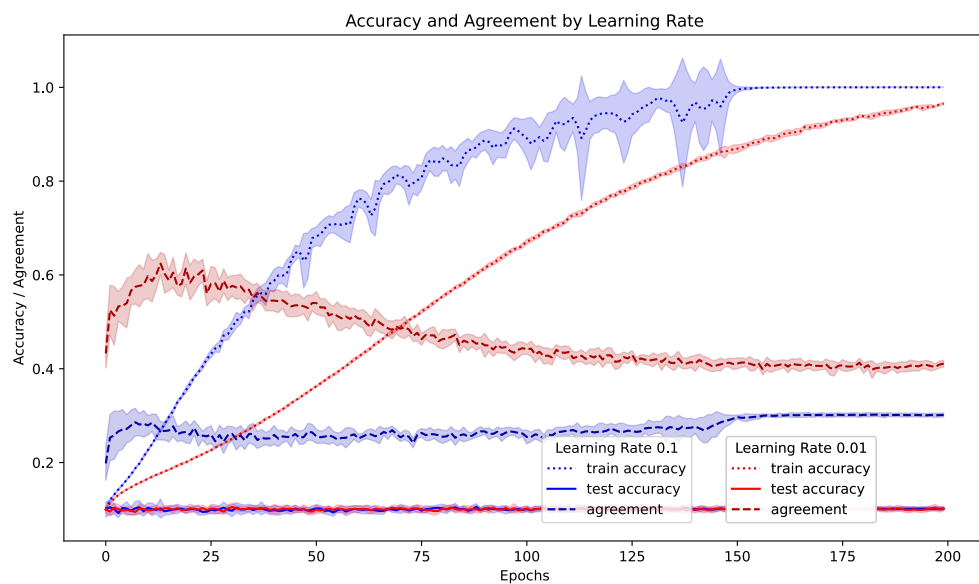


Figure 4: CIFAR10 with random labels

MNIST			FMNIST			CIFAR10			CIFAR10 - RAND		
#N	lr	Agree	#N	lr	Agree	#N	lr	Agree	#N	lr	Agree
256	0.1	99%	256	0.1	92%	256	0.1	62%	256	0.1	21%
256	0.01	99.5%	256	0.01	97%	256	0.01	71%	256	0.01	32%
512	0.1	99%	512	0.1	94%	512	0.1	67%	512	0.1	30%
512	0.01	99.5%	512	0.01	97%	512	0.01	80%	512	0.01	41%
1024	0.1	99%	1024	0.1	95%	1024	0.1	76%	1024	0.1	39%
1024	0.01	99.5%	1024	0.01	98%	1024	0.01	85%	1024	0.01	50%

Table 1: Agreement percentages across datasets with varying number of neurons in the hidden layer (#N) and learning rates (lr). The setup is the same as in [L.1](#) except for the number of training epochs, which is $\{50, 150, 150, 300\}$ for $\{\text{MNIST}, \text{FMNIST}, \text{CIFAR10}, \text{CIFAR10 random}\}$, respectively. In scenarios where agreement has not yet reached saturation, agreement is positively correlated with the width of the network.

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